

Title	Outage Probabilities of Orthogonal Multiple-Access Relaying Techniques with Imperfect Source-Relay Links
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Citation	IEEE Transactions on Wireless Communications
Issue Date	2014-12-18
Type	Journal Article
Text version	author
URL	http://hdl.handle.net/10119/12330
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Description	



Outage Probabilities of Orthogonal Multiple-Access Relaying Techniques with Imperfect Source-Relay Links

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Abstract—An outage probability that is independent of signaling schemes is theoretically derived in this paper for an orthogonal multiple-access relay channel (MARC) system, where the estimates of the information sequences sent from source nodes, regardless of whether or not they are correctly decoded at the relay, are exclusive-OR (XOR)-network-coded and forwarded by the relay to the destination. The MARC system described above is referred to as estimates-exploiting MARC (e-MARC) in this paper for convenience. Following the probability derivation of e-MARC, comparisons are then made with the outage probability of the orthogonal MARC with the Select Decode-and-Forward relaying strategy (MARC-SDF). It is found through simulations that when one of the source nodes is far away from both the relay and the destination, the e-MARC system is superior to MARC-SDF in terms of outage performance. We further numerically calculate the outage probabilities for two special cases, and compare them with the probability of e-MARC. Furthermore, the impact of the source correlation on the outage probability of the e-MARC system is also investigated.

Index Terms—multiple access relay channel, outage probability, decode-and-forward, Slepian-Wolf theorem, source-channel separation theorem.

I. INTRODUCTION

COOPERATIVE wireless networks have attracted a lot of attention of the wireless communication research community recently, since they provide a lot of design flexibility in the form of, e.g., coverage extension, and a variety of methodologies for multiple users quality-of-service (QoS) management. One of the cooperative wireless networks is multiple access relay [1], which consists of multiple source nodes, one relay and one common destination. The role of the relay is to assist the source nodes to improve the probability of successful transmission to the destination.

In multiple access relay channel (MARC) systems, decoded-and-forward (DF) is one of the most popular relaying strategies, and many DF-based excellent joint network-and-channel coding (JNCC) techniques for MARC systems have been

developed [2]–[7]. Most of the DF strategies used at the relay discard the information sequence estimates containing errors at the relay. This is because forwarding the erroneous estimates results in error propagation in the decoding process at the destination. As mentioned in [8], [9], performance and diversity gain are dramatically degraded due to the error propagation.

However, in practice, the source nodes are not geographically always close enough to the relay, and the average received signal-to-noise power ratio (SNR) is determined by the pathloss. Furthermore, the signal/signals received by the nodes may suffer from deep fade, depending on their locations. Error occurring in the source-relay links may well be eliminated by using powerful, capacity-achieving channel codes in the signaling scheme. Nevertheless, in many cases, the use of the powerful, close-capacity achieving codes per-link does not always provide reasonable solutions. This is because large latency is induced due to the iterative decoding process; moreover, the battery and memory resources of the relay may be limited to perform complicated decoding algorithms.

Quite recently, several excellent techniques have been proposed to exploit the erroneous estimates obtained at the relay [9]–[20]. In [16], the erroneous estimates of the information sent from the source nodes are, instead of discarding them, exclusive-OR (XOR)-network-coded at the relay, and forwarded to the destination to help reconstruct the information sequences; this MARC system is referred to as estimates-exploiting MARC (e-MARC) in this paper. The erroneous estimates in [16] are exploited in the decoding process at the destination by utilizing the knowledge of the bit-flipping probability p_{nc} between the two sequences: one is the network-coded sequence sent from the relay, and the other is its corresponding XOR-ed information sequence. The bit-flipping probability p_{nc} is referred to as *network correlation* in this paper.

Unlike techniques for the outage probability analysis [8], [21]–[24] for the cases where the relay stops forwarding if the estimates are found to contain error(s) at the relay, this paper analyzes the outage probability of e-MARC where erroneous estimates at the relay are further forwarded. The goal of this paper is twofold: (i) to re-formulate the e-MARC system according to the framework of the Slepian-Wolf theorem [25] for correlated source coding with a helper and analyze the e-MARC system's admissible rate region according to the re-formulation, and (ii) to derive the outage probability of the e-MARC system where all five links in the system (two

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This work was supported in part by the European Unions FP7 project, ICT-619555 RESCUE (Links-on-the-fly Technology for Robust, Efficient and Smart Communication in Unpredictable Environments), and also in part by the Japanese government funding program, Grant-in-Aid for Scientific Research(B), No. 23360170.

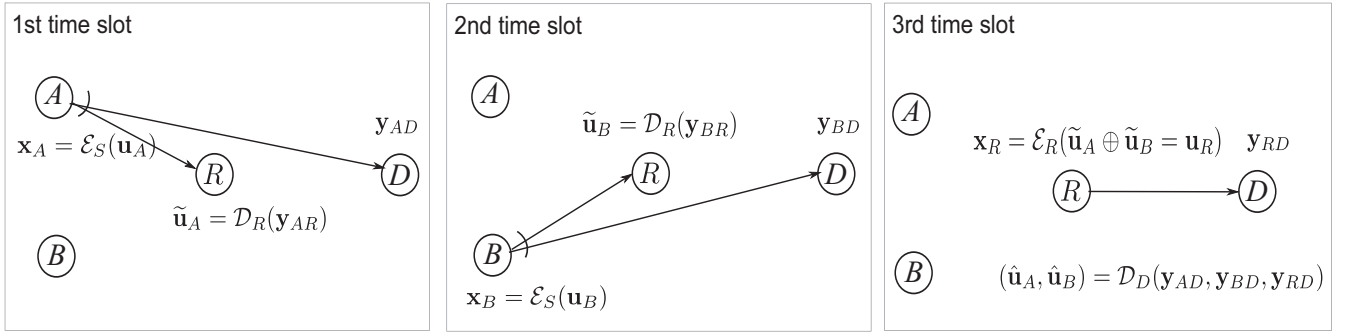


Fig. 1. Orthogonal e-MARC system model, where there are three time slots in a transmission cycle.

source-destination links, two source-relay links and one relay-destination link) are suffering from statistically independent block Rayleigh fading.

The theoretical analyses are based on the techniques presented in [26], [27]. However, in [26], the source-relay link(s) is assumed to be static, while in [27], they are assumed to be perfect, of which assumptions are not practical in real scenarios. In this paper, fading is taken into account in the source-relay link, and thus the source-relay link's error probability becomes a random variable, which makes the analysis for the admissible rate region to be challenging. Furthermore, to be able to compare with the theoretical outage performances of other relaying strategies applied in MARC, a theoretical limit of the source-relay link's error probability is found by utilizing the rate-distortion and inverse capacity functions in this paper. Then, it is shown that the outage probability of the e-MARC system, independent of signaling schemes, can be theoretically derived by a fivefold-integral over the admissible rate region with respect to the probability density functions (pdfs) of the five links' instantaneous SNRs.

The e-MARC system is then compared with network-coding-based orthogonal MARC systems (MARC-NC) [8] and orthogonal MARC with the select decode-and-forward relaying strategy (MARC-SDF) [24] in terms of the outage performance. It is observed through simulations that e-MARC is superior to MARC-NC. Furthermore, in Asymmetric scenario where one of the source nodes is far away from both the relay and the destination, e-MARC achieves better performance than MARC-SDF.

In addition, the process for deriving the probability is then applied to the following two special cases.

1) Source-relay links are assumed to be binary-symmetric channels (BSCs), and

2) The e-MARC system with practical signaling schemes presented in [16], where the relay only *extracts* (EX) the estimates output from differential detector (DD). The e-MARC system where the relay applies the detection strategy mentioned above is referred to as DDEX e-MARC system in this paper.

The fivefold-integral needed to obtain the outage probability can be reduced to simpler expressions, corresponding to their error probabilities of source-relay links. Numerical results show that theoretical outage probabilities with the DDEX e-MARC system are roughly 3.5 and 4.5 dB away from the probabilities of the e-MARC system independent of

signaling schemes in Symmetric and Asymmetric scenarios, respectively. This performance loss is due to the reason that we aim to reduce the computational complexity.

Finally, this paper investigates the impact of the correlation between the two source nodes on the outage probability of the e-MARC system. It is found that for the case two source nodes are highly correlated, the outage performance can still be improved by exploiting the *source correlation* even if only one of the source-destination links is reliable.

The rest of this paper is organized as follows: Section II introduces the system model of the e-MARC system assumed in this paper. The outage probability for e-MARC systems is theoretically derived in Section III. Numerical results of the outage probabilities for the universal case and two special cases are presented in Section IV. Furthermore, the results of simulations are presented to evaluate the performance of the scheme presented in [16] in terms of frame error rate (FER) in Section IV. Also, the theoretical outage probabilities of MARC-NC and that of MARC-SDF are included in Section IV for comparisons. The impact of the source correlation and network correlation on the outage probability is discussed in Section V. Finally, Section VI concludes the paper.

II. SYSTEM MODEL

Fig. 1 illustrates a basic model of the orthogonal e-MARC system assumed in this paper, where there are two source nodes A and B , one common relay R , and one common destination D . The K -bit length independent identically distributed (i.i.d.) binary information sequences generated from nodes A and B are denoted as $\mathbf{u}_A = \{u_A(k)\}_{k=1}^K$ and $\mathbf{u}_B = \{u_B(k)\}_{k=1}^K$, respectively. The signaling scheme used at the source node is denoted as $\mathcal{E}_S(\cdot)$, which consists of a serial concatenation of channel encoding and modulation. There are three time slots in one transmission cycle. In the first two time slots, nodes A and B respectively broadcast their coded sequences $\mathbf{x}_A = \mathcal{E}_S(\mathbf{u}_A) = \{x_A(m)\}_{m=1}^M$ and $\mathbf{x}_B = \mathcal{E}_S(\mathbf{u}_B) = \{x_B(m)\}_{m=1}^M$ to the relay R and destination D , and their corresponding received signals obtained at R and D are respectively written as

$$\begin{aligned} \mathbf{y}_{iR} &= h_{iR} \cdot \mathbf{x}_i + \mathbf{n}_{iR} \\ \mathbf{y}_{iD} &= h_{iD} \cdot \mathbf{x}_i + \mathbf{n}_{iD}, \quad i \in \{A, B\} \end{aligned} \quad (1)$$

where h_{iR} and h_{iD} indicate the channel coefficients of iR and iD links with the source node i , respectively. \mathbf{n}_{iR} and

TABLE I
NOTATIONS.

i	Source node, $i \in \{A, B\}$
R	Relay
D	Destination
\mathbf{u}_i	Information sequence of i
$\tilde{\mathbf{u}}_i$	Estimates of \mathbf{u}_i received at R
$\hat{\mathbf{u}}_i$	Estimates of \mathbf{u}_i received at D
\mathbf{x}_i	Coded sequences of node i
\mathbf{y}_{iR}	Received signals at R sent from i
\mathbf{y}_{iD}	Received signals at D sent from i
\mathbf{y}_{RD}	Received signals at D sent from R
p_i	Error probability/ rate of iR link
p_R	Error probability/ rate of RD link
\hat{p}_i	Theoretical limit of p_i
\hat{p}_R	Theoretical limit of p_R
\tilde{p}_i	Non zero value of \hat{p}_i
\tilde{p}_R	Non zero value of \hat{p}_R
γ_{iR}	Instantaneous SNR of iR link
γ_{iD}	Instantaneous SNR of iD link
γ_{RD}	Instantaneous SNR of RD link
γ_i^*	$p_i = 0$ if $\gamma_{iR} > \gamma_i^*$
γ_R^*	$p_R = 0$ if $\gamma_{RD} > \gamma_R^*$
\mathcal{E}_S	Signaling scheme at i
\mathcal{E}_R	Signaling scheme at R
\mathcal{R}_i	Source coding rate at i
\mathcal{R}_R	Source coding rate at R
\mathcal{R}_{ci}	Spectrum efficiency of \mathcal{E}_S
\mathcal{R}_{cR}	Spectrum efficiency of \mathcal{E}_R
$H_b(\cdot)$	Binary entropy function
$*$	convolution, $\alpha * \beta = \alpha(1 - \beta) + \beta(1 - \alpha)$
δ	$H_b(p_A * p_B * p_R)$
$\hat{\delta}$	Theoretical limit of δ
$\check{\delta}$	δ with constant p_i
p_s	Source correlation
p_{nc}	Network correlation; $p_{nc} = p_A * p_B$

\mathbf{n}_{iD} indicate the vectors of independent zero-mean complex additive white Gaussian noise (AWGN) of the iR and iD links, respectively, with variance $\sigma_{iR}^2 = \sigma_{iD}^2 = \sigma^2$ per dimension. The iR and iD links are also referred to as the intra and direct links, respectively, in this paper.

The receiver applied at the relay R is denoted as $\mathcal{D}_R(\cdot)$, composed of signal detection and decoding, which corresponds to the inverse structure of $\mathcal{E}_S(\cdot)$. The estimates $\tilde{\mathbf{u}}_i = \mathcal{D}_R(\mathbf{y}_{iR})$ of \mathbf{u}_i obtained at R may contain errors due to the variation of the iR link. The error rate of the iR link is represented by

$$p_i = \mathfrak{B}(\mathbf{u}_i, \tilde{\mathbf{u}}_i) = \frac{\sum_{k=1}^K |u_i(k) - \tilde{u}_i(k)|}{K}, \quad i \in \{A, B\}. \quad (2)$$

In the e-MARC system model, the estimates $\tilde{\mathbf{u}}_A$ and $\tilde{\mathbf{u}}_B$, are always joint network-channel coded at R regardless of whether they are correct or not, as

$$\mathbf{x}_R = \mathcal{E}_R(\mathbf{u}_R) = \mathcal{E}_R(\tilde{\mathbf{u}}_A \oplus \tilde{\mathbf{u}}_B), \quad (3)$$

where the notation \oplus denotes a binary exclusive-OR (XOR) operation and $\mathcal{E}_R(\cdot)$ represents the signaling scheme applied at R , including channel encoding and modulation. The destination D obtains the signal vector \mathbf{y}_{RD} of \mathbf{x}_R sent via the relay-destination (RD) link as

$$\mathbf{y}_{RD} = h_{RD} \cdot \mathbf{x}_R + \mathbf{n}_{RD}, \quad (4)$$

where h_{RD} and \mathbf{n}_{RD} indicate the channel coefficient and AWGN vector of the RD link with variance $\sigma_{RD}^2 = \sigma^2$, respectively. The estimates of \mathbf{u}_R obtained at the destination

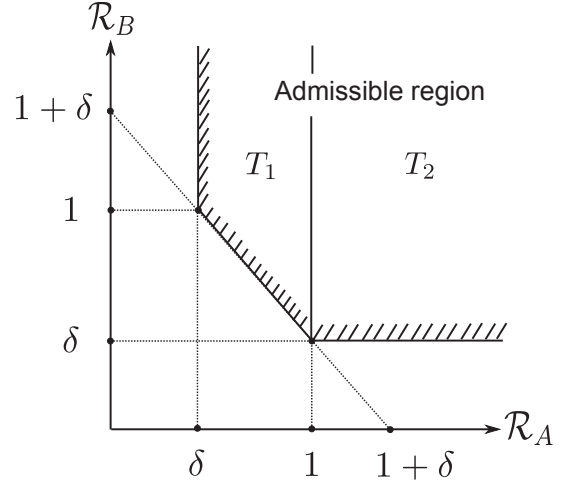


Fig. 2. Admissible rate region of the rate pair $(\mathcal{R}_A, \mathcal{R}_B)$ given $\mathcal{R}_R \geq I(\mathbf{u}_R; \hat{\mathbf{u}}_R)$ for lossless compression, where $\delta = H_b(p_A * p_B * p_R)$.

is denoted as $\hat{\mathbf{u}}_R$, and the error rate of the RD link is $p_R = \mathfrak{B}(\mathbf{u}_R, \hat{\mathbf{u}}_R)$. Finally, to obtain the estimates $\hat{\mathbf{u}}_A$ and $\hat{\mathbf{u}}_B$ of the information sequences \mathbf{u}_A and \mathbf{u}_B , respectively, at the destination, decoding of JNCC is performed on the received signal vectors \mathbf{y}_{AD} and \mathbf{y}_{BD} with the help of the signal vector \mathbf{y}_{RD} .

All the links in this paper are assumed to suffer from block Rayleigh fading, with which h_{iR} , h_{iD} and h_{RD} are assumed to be constant over one coded sequence but vary independently transmission-by-transmission and link-by-link. Without loss of generality, we assume that $E[|h_{iR}|^2] = E[|h_{iD}|^2] = E[|h_{RD}|^2] = 1$. The error rates p_i and p_R , thus, vary in each transmission cycle. The instantaneous SNRs γ_{iR} , γ_{iD} and γ_{RD} of the links are then given by

$$\begin{aligned} \gamma_{iR} &= |h_{iR}|^2 \cdot \Gamma_{iR} \\ \gamma_{iD} &= |h_{iD}|^2 \cdot \Gamma_{iD} \\ \gamma_{RD} &= |h_{RD}|^2 \cdot \Gamma_{RD} \end{aligned} \quad (5)$$

where Γ_{iR} , Γ_{iD} and Γ_{RD} represent the average SNRs of the intra, direct and RD links, respectively.

III. DERIVATION FOR OUTAGE PROBABILITY OF E-MARC

A. Outage event for each transmission cycle

The outage probability of the e-MARC system is derived in this section. The e-MARC system model shown in Fig. 1 can be viewed as a system having two source nodes A and B , and one helper R . Letting \mathcal{R}_A and \mathcal{R}_B be the source coding rate of A and B , respectively. To investigate the admissible rate region for the two independent source nodes with one helper, we thus invoke the theorem [28, Theorem 10.4.], according to which the information sequences \mathbf{u}_A and \mathbf{u}_B can be successfully recovered at the destination if

$$\begin{aligned}
\mathcal{R}_A &\geq H(\mathbf{u}_A|\mathbf{u}_B, \hat{\mathbf{u}}_R), \\
\mathcal{R}_B &\geq H(\mathbf{u}_B|\mathbf{u}_A, \hat{\mathbf{u}}_R), \\
\mathcal{R}_A + \mathcal{R}_B &\geq H(\mathbf{u}_A, \mathbf{u}_B|\hat{\mathbf{u}}_R), \\
\mathcal{R}_R &\geq I(\mathbf{u}_R; \hat{\mathbf{u}}_R),
\end{aligned} \tag{6}$$

where \mathcal{R}_R represents the source coding rate at R , and the coded side information provided by the relay R (helper) helps reduce the rate \mathcal{R}_A and \mathcal{R}_B . Then, by using the chain rule, (6) can be re-formulated (see Appendix A) as follows

$$\mathcal{R}_A \geq \delta, \tag{7a}$$

$$\mathcal{R}_B \geq \delta, \tag{7b}$$

$$\mathcal{R}_A + \mathcal{R}_B \geq \delta + 1, \tag{7c}$$

$$\mathcal{R}_R \geq 1 - H_b(p_R), \tag{7d}$$

where

$$\delta = H_b(p_A * p_B * p_R). \tag{8}$$

The operation $\alpha * \beta$ is defined as $\alpha(1 - \beta) + \beta(1 - \alpha)$, and $H_b(\cdot)$ denotes the binary entropy function [29]. To simplify the derivation, we first assume that the inequality (7d) is satisfied. With a given \mathcal{R}_R , the admissible rate region of $(\mathcal{R}_A, \mathcal{R}_B)$ for each transmission cycle is obtained, as shown in Fig. 2.

For a transmission cycle, the outage event is defined as: one or both of the information sequences \mathbf{u}_A and \mathbf{u}_B cannot be successfully recovered at the destination. Therefore, given a value of \mathcal{R}_R , the outage event happens when the pair $(\mathcal{R}_A, \mathcal{R}_B)$ falls outside the admissible rate region. As shown in Fig. 2, the entire admissible rate region can be divided into two parts T_1 and T_2 . ε_1 and ε_2 denote the events that the rate pair $(\mathcal{R}_A, \mathcal{R}_B)$ falls into T_1 and T_2 , respectively, as

$$\begin{aligned}
\varepsilon_1 &= \{(\mathcal{R}_A, \mathcal{R}_B) \in T_1\} \\
&= \{\delta \leq \mathcal{R}_A \leq 1\} \wedge \{\mathcal{R}_A + \mathcal{R}_B \geq \delta + 1\}, \\
\varepsilon_2 &= \{(\mathcal{R}_A, \mathcal{R}_B) \in T_2\} = \{\mathcal{R}_A \geq 1\} \wedge \{\mathcal{R}_B \geq \delta\}.
\end{aligned} \tag{9}$$

where the symbols ‘ \vee ’ and ‘ \wedge ’ denote as the logical ‘or’ and ‘and’ operators, respectively. Therefore, with the assumption that (7d) is satisfied, the outage event of the e-MARC system for each transmission cycle is obtained, as

$$\text{OUT} = \overline{\{\varepsilon_1 \vee \varepsilon_2\}}. \tag{10}$$

where the $\overline{\{\varepsilon_1 \vee \varepsilon_2\}}$ denotes the complement of event $\{\varepsilon_1 \vee \varepsilon_2\}$.

According to Shannon’s source-channel separation theorem, the relationship between the instantaneous SNR γ_{iD} and its corresponding source coding rate \mathcal{R}_i is given by [30]

$$\mathcal{R}_i = \frac{C(\gamma_{iD})}{\mathcal{R}_{ci}} = f_i(\gamma_{iD}), \quad i \in \{A, B\} \tag{11}$$

where it is assumed that a capacity-achieving channel code is used in iD link. Here \mathcal{R}_{ci} represents the spectrum efficiency of the signaling scheme $\mathcal{E}_S(\cdot)$, including the channel coding scheme and the modulation multiplicity, and $C(\alpha) = \log_2(1 +$

$\alpha)$. It should be emphasized that the optimality of source-channel separation holds for Rayleigh fading MARC systems where all the links are orthogonal [31], [32].

Since the function $f_i(\cdot)$ is one-to-one mapping, the events ε_1 and ε_2 can be, respectively, expressed as

$$\begin{aligned}
\varepsilon_1 &= \{f_A^{-1}(\delta) \leq \gamma_{AD} \leq f_A^{-1}(1)\} \wedge \{f_B^{-1}(\omega) \leq \gamma_{BD}\} \\
\varepsilon_2 &= \{f_A^{-1}(1) \leq \gamma_{AD}\} \wedge \{f_B^{-1}(\delta) \leq \gamma_{BD}\},
\end{aligned} \tag{12}$$

where

$$\omega = \delta + 1 - f_A(\gamma_{AD}). \tag{13}$$

B. Theoretical limits of p_A and p_B

The boundaries of the events ε_1 and ε_2 in (12) are involved with the iR link error probability¹ p_i , and the value of the p_i is a function of γ_{iR} . However, the relationship between p_i and γ_{iR} depends on the signaling schemes employed in the iR link. Furthermore, it is quite common that p_i cannot be explicitly expressed as a function of γ_{iR} , if specific channel coding or modulation schemes are used. Therefore, instead, we aim to derive a theoretical limit for the value of p_i , given a γ_{iR} value in the following.

According to Shannon’s lossy source-channel separation theorem [33], the information sequence \mathbf{u}_i can be transmitted over the iR link with a distortion level \mathcal{D}_i if

$$\mathcal{R}_i(\mathcal{D}_i)\mathcal{R}_{ci} \leq C(\gamma_{iR}), \quad i \in \{A, B\}. \tag{14}$$

With the Hamming distortion measure, the distortion \mathcal{D}_i is equivalent to the error probability p_i , and thus according to [29], the rate-distortion function $\mathcal{R}_i(\mathcal{D}_i)$ is represented as

$$\mathcal{R}_i(\mathcal{D}_i) = 1 - H_b(p_i). \tag{15}$$

By assuming that a capacity-achieving channel code is applied at the iR link, the equality in (14) holds, yielding the theoretical limit \hat{p}_i of p_i for any given γ_{iR} , as

$$\hat{p}_i(\gamma_{iR}) = \begin{cases} H_b^{-1}(1 - f_i(\gamma_{iR})) = \tilde{p}_i(\gamma_{iR}), & 0 \leq \gamma_{iR} < \gamma_i^* \\ 0, & \gamma_{iR} \geq \gamma_i^* \end{cases} \tag{16}$$

where $H_b^{-1}(\cdot)$ denotes the inverse function of $H_b(\cdot)$, and γ_i^* is a threshold value of the instantaneous SNR γ_{iR} such that $f_i(\gamma_i^*) = 1$.

C. Theoretical limit of p_R

In this subsection, the assumption that the inequality (7d) holds in (12) is eliminated, such that the variation of the rate \mathcal{R}_R can be taken into account when deriving the outage probability in the next subsection. In the same way as deriving (11), the relationship between the instantaneous SNR γ_{RD} and rate \mathcal{R}_R can be expressed, as

¹The length of information sequence K is assumed to be infinite for deriving the fully theoretical outage probability.

$$\mathcal{R}_R = \frac{C(\gamma_{RD})}{\mathcal{R}_{cR}} = f_R(\gamma_{RD}) \geq 1 - H_b(p_R) \quad (17)$$

where \mathcal{R}_{cR} represents the spectrum efficiency of the signaling scheme $\mathcal{E}_R(\cdot)$. Considering the constraint imposed on \mathcal{R}_R (i.e., $\mathcal{R}_R \geq 1 - H_b(p_R)$), it is found from (17) that, although the variation range of the rate \mathcal{R}_R is $[0, \infty)$, p_R is reduced to zero when the value of γ_{RD} is larger than $f_R^{-1}(1)$. Therefore, by (17), the theoretical limit \hat{p}_R of p_R for any given instantaneous SNR value γ_{RD} , is

$$\hat{p}_R(\gamma_{RD}) = \begin{cases} H_b^{-1}(1 - f_R(\gamma_{RD})) = \tilde{p}_R(\gamma_{RD}), & 0 \leq \gamma_{RD} < \gamma_R^* \\ 0, & \gamma_{RD} \geq \gamma_R^* \end{cases} \quad (18)$$

where $\gamma_R^* = f_R^{-1}(1)$.

By replacing p_i and p_R with $\hat{p}_i(\gamma_{iR})$ and $\hat{p}_R(\gamma_{RD})$ in (12), respectively, the boundaries of events ε_1 and ε_2 in (12) are converted to the domains of instantaneous SNR, as

$$\begin{aligned} \varepsilon_1 &= \{f_A^{-1}(\hat{\delta}) \leq \gamma_{AD} \leq f_A^{-1}(1)\} \wedge \{f_B^{-1}(\hat{\omega}) \leq \gamma_{BD}\} \\ \varepsilon_2 &= \{f_A^{-1}(1) \leq \gamma_{AD}\} \wedge \{f_B^{-1}(\hat{\delta}) \leq \gamma_{BD}\}, \end{aligned} \quad (19)$$

where

$$\begin{aligned} \hat{\delta} &= H_b(\hat{p}_A(\gamma_{AR}) * \hat{p}_B(\gamma_{BR}) * \hat{p}_R(\gamma_{RD})) \\ \hat{\omega} &= \hat{\delta} + 1 - f_A(\gamma_{AD}). \end{aligned} \quad (20)$$

D. Outage probability of e-MARC

Recall that in Section II all the links are assumed to be statistically independent and their corresponding instantaneous SNRs are Rayleigh distributed,² transmission-by-transmission and link-by-link. Hence, the probabilities of ε_1 and ε_2 are calculated, as

$$\begin{aligned} \Pr(\varepsilon_1) &= \Pr\left(\{f_A^{-1}(\hat{\delta}) \leq \gamma_{AD} \leq f_A^{-1}(1)\} \wedge \{f_B^{-1}(\hat{\omega}) \leq \gamma_{BD}\}\right) \\ &= \underbrace{\iiint_V \int_{f_A^{-1}(\hat{\delta})}^{f_A^{-1}(1)} p(\gamma_{AD}) d\gamma_{AD} \int_{f_B^{-1}(\hat{\omega})}^{\infty} p(\gamma_{BD}) d\gamma_{BD}}_{g_1} \\ &\quad p(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) d\gamma_{AR} d\gamma_{BR} d\gamma_{RD} \\ &= \underbrace{\iiint_V \frac{1}{\Gamma_{AD}} \int_{f_A^{-1}(\hat{\delta})}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega})}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}}\right) d\gamma_{AD}}_{g_1} \\ &\quad p(\gamma_{AR}) p(\gamma_{BR}) p(\gamma_{RD}) d\gamma_{AR} d\gamma_{BR} d\gamma_{RD} \\ &= \mathcal{J}[g_1; V] \end{aligned} \quad (21)$$

² $p(\gamma_q) = \frac{1}{\Gamma_q} \exp\left(-\frac{\gamma_q}{\Gamma_q}\right)$, $q \in \{AR, BR, AD, BD, RD\}$

and

$$\begin{aligned} \Pr(\varepsilon_2) &= \Pr\left(\{f_A^{-1}(1) \leq \gamma_{AD}\} \wedge \{f_B^{-1}(\hat{\delta}) \leq \gamma_{BD}\}\right) \\ &= \underbrace{\iiint_V \int_{f_A^{-1}(1)}^{\infty} p(\gamma_{AD}) d\gamma_{AD} \int_{f_B^{-1}(\hat{\delta})}^{\infty} p(\gamma_{BD}) d\gamma_{BD}}_{g_2} \\ &\quad p(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) d\gamma_{AR} d\gamma_{BR} d\gamma_{RD} \\ &= \underbrace{\iiint_V \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{f_B^{-1}(\hat{\delta})}{\Gamma_{BD}}\right)}_{g_2} \\ &\quad p(\gamma_{AR}) p(\gamma_{BR}) p(\gamma_{RD}) d\gamma_{AR} d\gamma_{BR} d\gamma_{RD} \\ &= \mathcal{J}[g_2; V], \end{aligned} \quad (22)$$

where the domain of the threefold integral is

$$\begin{aligned} V &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{R}^+, \gamma_{BR} \in \mathbb{R}^+, \gamma_{RD} \in \mathbb{R}^+\}, \\ \mathbb{R}^+ &= [0, \infty). \end{aligned} \quad (23)$$

Finally, the outage probability of the e-MARC system can be obtained, as

$$P_{\text{out}} = 1 - (\Pr(\varepsilon_1) + \Pr(\varepsilon_2)). \quad (24)$$

It may be difficult to calculate the integrals shown in (21) and (22) in closed form. Hence, the results of (21) and (22) in this paper are numerically obtained by using functions provided in [34]. Moreover, (21) and (22) can be respectively divided into eight sub-integrals according to different domains, which makes the numerical calculation of (21) and (22) tractable. The divisions of (21) and (22) are listed in Appendix B.

IV. NUMERICAL RESULTS

A. Outage probability of e-MARC

Fig. 3 shows the numerical results of the outage probability for the e-MARC system that are independent of signaling schemes, obtained by calculating from (24) with (21) and (22) shown in Section III, where $\Gamma_{AD} = \Gamma_{BD}$ and $\Gamma_{RD} = \Gamma_{AD} + 3$ dB is assumed. The values of \mathcal{R}_{ci} and \mathcal{R}_{cR} are set to 1/2. The curve of the probability with perfect intra links is obtained from [27]. It is found that as the quality of intra links degrades, the gap increases between the probabilities with the perfect and with the imperfect intra links.

B. Outage performance comparisons

To evaluate the impact of the utilization of the erroneous estimates received at the relay, we also compare the outage probability of the e-MARC system with the theoretical outage probabilities with MARC-NC [8] and MARC-SDF [24].

TABLE II
SETTINGS OF SYMMETRIC AND ASYMMETRIC SCENARIOS.

Scenario	Γ_{AD}	Γ_{BD}	Γ_{AR}	Γ_{BR}	Γ_{RD}
Symmetric	X	X	$X+\Delta$	$X+\Delta$	$X+\Delta$
Asymmetric	X	$X-L$	$X+\Delta$	$X+\Delta-L$	$X+\Delta$

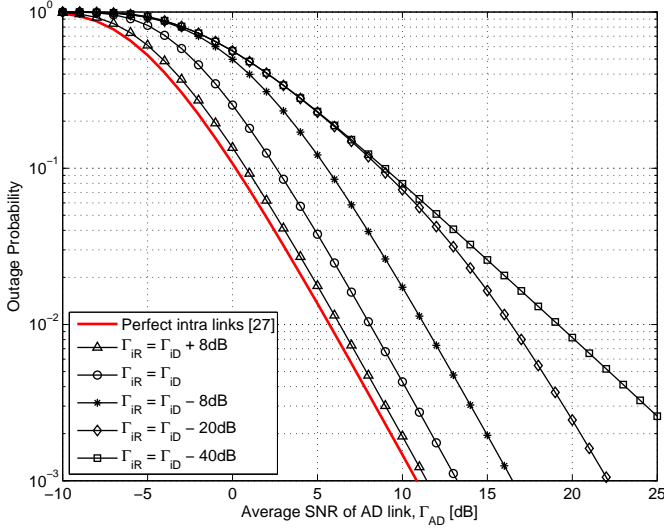


Fig. 3. Outage probabilities of e-MARC that are independent of signaling schemes, where $\Gamma_{AD} = \Gamma_{BD}$ and $\Gamma_{RD} = \Gamma_{AD} + 3$ dB is assumed, and $\mathcal{R}_{ci} = \mathcal{R}_{cr} = 1/2$.

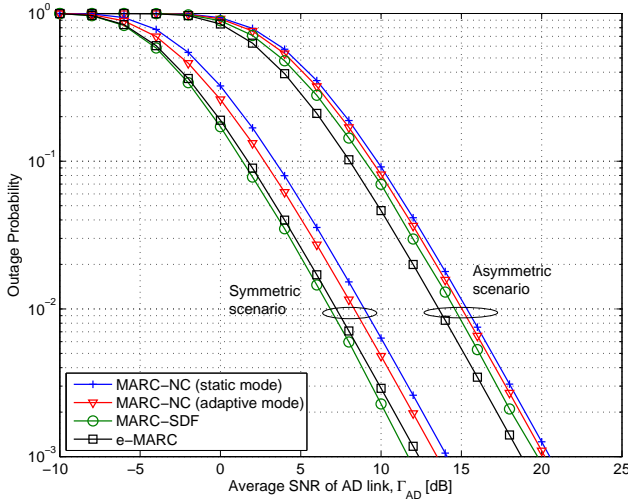


Fig. 4. Outage probabilities of e-MARC are in comparisons with those with MARC-NC [8], and with MARC-SDF [24]. $\mathcal{R}_{ci} = \mathcal{R}_{cr} = 1/2$.

Results of the outage probability analysis in Symmetric and Asymmetric scenarios are demonstrated in Fig. 4 for comparisons. The scenarios settings exemplifying the symmetric and asymmetric cases are summarized in Table II, where X represents the average SNR of the AD link in dB; Δ and L denoting additional gain and loss due to the shorter and longer distance, respectively, and are set to 3 dB and 10 dB in this subsection.

The theoretical outage probabilities of MARC-NC, analyzed in a separate network-and-channel coding (SNCC) framework, were included in Fig. 4. The outage probabilities³ in static and in adaptive mode are obtained by [8, eqs. (4)-(5)] and [8, eqs. (4),(8)-(9),(11)], respectively. It is found that the e-MARC system outperforms MARC-NC in terms of outage performance.

In principle, as stated in [24], a header (i.e., flag) is

³The probabilities are computed at a system level [8].

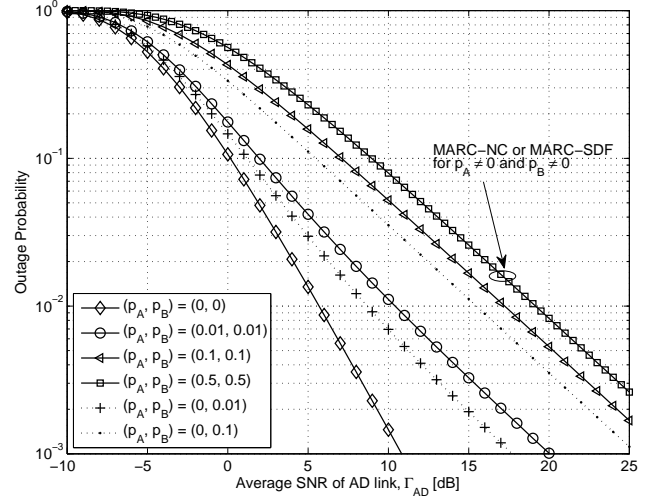


Fig. 5. Theoretical outage probabilities of e-MARC with static intra links, where $\Gamma_{AD} = \Gamma_{BD}$ and $\Gamma_{RD} = \Gamma_{AD} + 3$ dB. $\mathcal{R}_{ci} = \mathcal{R}_{cr} = 1/2$

used at the relay in MARC-SDF to identify the correctly decoded source nodes, and furthermore, the header has to be protected with a very powerful error correction code. On the other hand, the relay in e-MARC always performs network-coding and does not add header to the forwarded XOR-ed sequence. Therefore, it is not reasonable to make comparisons between MARC-SDF and e-MARC, because MARC-SDF and e-MARC belong to different categories of MARC (i.e., MARC-SDF is address-based, while e-MARC is non-address-based⁴).

Up to our best knowledge, there is no explicit mathematical expressions to calculate the theoretical outage probability of MARC-SDF. Nevertheless, it is still meaningful to include MARC-SDF performance curves as a reference. The performance results with MARC-SDF, shown in Fig. 4, are all based on the Monte-Carlo method according to [24, eqs. (4)-(6)] with a modification of [36, eq. (12)], while all the curves for e-MARC are the theoretical results.

Deriving explicit mathematical expression of the outage probability with MARC-SDF is out of the scope of this paper; hence, we emphasize again that the conclusions for the superiority/inferiority of e-MARC related to MARC-SDF in the Asymmetric/Symmetric scenarios are based on simulations. Rigorous mathematical analyzes for the results shown in Fig. 4 are left as future study.

C. Special cases

1) *Static intra links*: In the case that intra links are static (not suffering from fading), AR and BR links can be modeled as BSCs with constant crossover probabilities p_A and p_B , respectively. This scenario may be exemplified by sensor networks where only the destination moves. In that case, averaging the AR and BR intra link variations over their pdfs can be eliminated from (21) and (22). Accordingly, the derivations of $\Pr(\varepsilon_1)$ and $\Pr(\varepsilon_2)$ can be reduced to

⁴In practical applications, the value of $p_A * p_B$ can be directly estimated and exploited at the destination without requiring any overhead transmission from the relay [35, Section III-B].

$$\Pr(\varepsilon_1) = \frac{1}{\Gamma_{AD}} \int_0^\infty \int_{f_A^{-1}(\delta)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\dot{\omega})}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}}\right) d\gamma_{AD} p(\gamma_{FR})$$

$$\Pr(\varepsilon_2) = \int_0^\infty \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{f_B^{-1}(\delta)}{\Gamma_{BD}}\right) p(\gamma_{RD}) d\gamma_{RD}, \quad (25)$$

where

$$\ddot{\delta} = H_b(p_A * p_B * \hat{p}_R(\gamma_{RD}))$$

$$\dot{\omega} = \ddot{\delta} + 1 - f_A(\gamma_{AD}). \quad (26)$$

The results of the theoretical outage probabilities calculated in this special case are shown in Fig. 5. It can be clearly observed in Fig. 5 that the outage probability achieves the second-order diversity when $p_A = p_B = 0$, and matches the probability curve with perfect intra links in Fig. 3. When both intra links are imperfect but with relatively small crossover probabilities (e.g., when $p_A = p_B = 0.01$), it is found that the decay of the outage probability follows the second order diversity in the low Γ_{iD} regime. However, the diversity order asymptotically converges to the first order as Γ_{iD} increases. It should be noticed that for the case that $p_A \neq 0$ and $p_B \neq 0$, the relay remains silent with MARC-SDF or MARC-NC; however, with e-MARC, the erroneous estimates forwarded from the relay still contribute to reducing the outage probability.

In the case one of the intra links is perfect but the other one is imperfect, the helper's contribution, provided by the relay, depends on the quality of the imperfect link. This is because in (7a)-(7d) provided in Section III, the value of $p_A * p_B$ is a dominating factor that determines the size of the admissible rate region of $(\mathcal{R}_A, \mathcal{R}_B)$. Hence, p_A dominates the value of $p_A * p_B$ if $p_B = 0$, and vice versa.

2) *DDEX e-MARC system [16]*: The derivation for the outage probability of the e-MARC system shown in Section III can further be used to evaluate the efficiency of the DDEX e-MARC system proposed by [16] where the erroneous estimates are also utilized. In [16], the information sequences \mathbf{u}_i for $i \in \{A, B\}$ are first encoded by recursive systematic convolutional (RSC) codes, and then broadcasted from the source nodes using binary differential phase-shift-keying (DPSK) modulation.⁵ To achieve low computational complexity and small latency, instead of performing fully-iterative decoding between the decoders of the accumulator (ACC) [37] and RSC code, the relay R obtains the estimates $\tilde{\mathbf{u}}_i$ of \mathbf{u}_i by the DDEX detection strategy, where the relay simply extracts the systematic part output from the differential detector. The estimates $\tilde{\mathbf{u}}_i$ for $i \in \{A, B\}$ are detected in error with a high probability because no error correction is performed at R , in exchange for saving battery and memory resources due to low computational complexity. The obtained estimates are XOR-network-coded at R and forwarded to the destination D for improving the FER performance.

⁵The accumulator (ACC) [37] followed by coherent PSK is equivalent to DPSK.

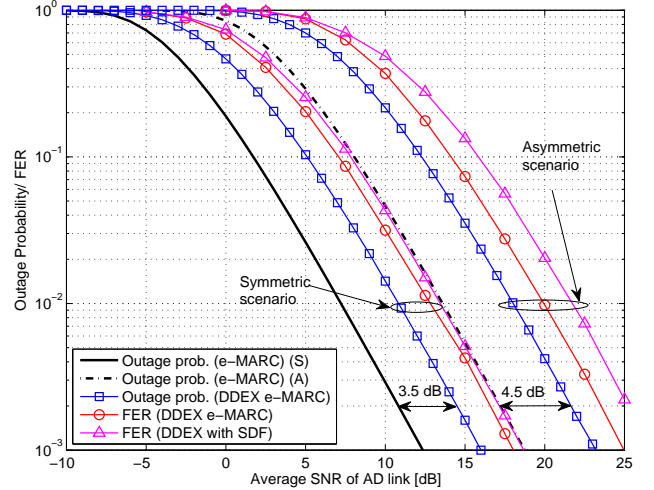


Fig. 6. Theoretical outage probabilities of DDEX e-MARC system [16]. Notations (S) and (A) denote Symmetric and Asymmetric scenarios listed in Table II, respectively. Select DF (SDF) strategy is same as that presented in [24].

TABLE III
PARAMETERS USED IN FER SIMULATIONS.

Length of information sequence (K)	2048 bits
Generator polynomial of RSC code	$(07, 05)_8$
Generator polynomial of ACC	$(03, 02)_8$
Signaling scheme \mathcal{E}_S	[16, Fig. 1]
Signaling scheme \mathcal{E}_R	[16, Fig. 1]
Spectrum efficiency of \mathcal{E}_S (K/M)	1/2
Spectrum efficiency of \mathcal{E}_R (K/M)	1/2
Receiver \mathcal{D}_R	DDEX
Receiver \mathcal{D}_D	[16, Fig. 4]
Decoding algorithm	log-MAP
Decoding of \mathcal{E}_S at D	10 iterations
Decoding between \mathcal{E}_S and \mathcal{E}_R at D	10 iterations
Interleavers	Random

As shown in [16], DPSK modulation and differential detection are used at the source node and relay, respectively, and hence according to [38], the error probability of the *extracted* estimates of the information sequence is expressed as

$$p_i(\gamma_{iR}) = \frac{1}{2} \exp(-\gamma_{iR}), \quad i \in \{A, B\}. \quad (27)$$

The theoretical outage probability for the DDEX e-MARC system (i.e., the theoretical achievable FER performance of the system), is obtained by replacing (16) with (27) in (24). Results of outage probabilities in Symmetric and Asymmetric scenarios are demonstrated in Fig. 6 for comparison. The settings of the both scenarios are same as those in the subsection IV-B.

In Symmetric scenario, as shown in Fig. 6, the outage probability of DDEX e-MARC system is roughly 3.5 dB away from the outage probability of e-MARC that is independent of the signaling scheme. The gap is due to the relatively high p_i value with DDEX. In Asymmetric scenario, the gap is further increased by 1 dB, resulting in a 4.5-dB gap from the outage probability of e-MARC. However, a large number of addition, multiplication and comparison operations can be eliminated on obtaining the estimates $\tilde{\mathbf{u}}_i$, according to [39, Section 4.1].

The FER performance results of the DDEX e-MARC system obtained through simulations are also included in Fig. 6. Here FER is defined as follows: number of the transmission cycles where either one or both of the information sequences sent from the nodes A and B cannot successfully recovered at D even with the help of R , divided by the total number of transmission cycles. The parameters used in the simulations are summarized in Table III. It is found from Fig. 6 that both in Symmetric and Asymmetric scenarios, there is a roughly 1.8-dB loss with the practical FER performance from the corresponding theoretical outage probabilities. The loss is because the theoretical outage probabilities are derived assuming the use of capacity-achieving codes in the the iD and RD links.

V. IMPACT OF CORRELATION

In this section, we further extended our analysis to the e-MARC system where the correlation between two source nodes is taken into account. The correlation of source nodes is modeled by random bit-flipping with the flipping probability $p_s = \mathfrak{B}(\mathbf{u}_A, \mathbf{u}_B)$. The admissible rate region for the e-MARC system with the correlation p_s between source nodes is analyzed in (40) in Appendix A. To investigate the impact of the source correlation p_s and the network correlation $p_{nc} = p_A * p_B = \mathfrak{B}(\mathbf{u}_A \oplus \mathbf{u}_B, \tilde{\mathbf{u}}_A \oplus \tilde{\mathbf{u}}_B)$, three extreme scenarios: $(\Gamma_{AD}, \Gamma_{BD}, p_R) = (30 \text{ dB}, -30 \text{ dB}, 0)$, $(30 \text{ dB}, -30 \text{ dB}, 0.5)$ and $(-30 \text{ dB}, -30 \text{ dB}, 0.5)$ are investigated, respectively. The values of \mathcal{R}_{ci} and \mathcal{R}_{cR} are set to $1/2$.

A. Source correlation p_s

We first focus on the impact of source correlation p_s on the e-MARC outage probability. For the scenarios $(\Gamma_{AD}, \Gamma_{BD}, p_R) = (30 \text{ dB}, -30 \text{ dB}, 0)$ and $(30 \text{ dB}, -30 \text{ dB}, 0.5)$, the soft channel values of the received signal vector \mathbf{y}_{AD} are reliable while the values of \mathbf{y}_{BD} are unreliable at the destination D . Even in this case, the information sequence \mathbf{u}_B can still be recovered with a high probability if two source nodes are highly correlated. This is because when \mathbf{u}_A and \mathbf{u}_B are highly correlated, reliable *a priori* information for the recovery of \mathbf{u}_B can be directly acquired from *extrinsic* information related to \mathbf{u}_A during the decoding process. Hence, it can be concluded that, with or without the relay, when at least one of the direct links is reliable and the two source nodes are highly correlated, a very low probability can be achieved. This situation is demonstrated in Figs. 7(a) and 7(b), where the probability is shown as a function of p_s and p_{nc} .

B. Network correlation p_{nc}

Then, let us focus on the impact of the network correlation p_{nc} . Assuming \mathbf{u}_A and \mathbf{u}_B are uncorrelated, but $\mathbf{u}_A \oplus \mathbf{u}_B$ and $\tilde{\mathbf{u}}_A \oplus \tilde{\mathbf{u}}_B$ are fully correlated (i.e., $\mathbf{u}_A \oplus \mathbf{u}_B = \tilde{\mathbf{u}}_A \oplus \tilde{\mathbf{u}}_B$ and $p_{nc} = 0$). For the scenario $(\Gamma_{AD}, \Gamma_{BD}, p_R) = (30 \text{ dB}, -30 \text{ dB}, 0)$, the soft channel values of \mathbf{y}_{AD} and \mathbf{y}_{RD} at the destination are reliable while the values of \mathbf{y}_{BD} are unreliable. In this scenario, \mathbf{u}_A and \mathbf{u}_R can be recovered with a high probability after the decoding of \mathbf{y}_{AD} and \mathbf{y}_{RD} , respectively.

\mathbf{u}_B can also be recovered with the aid of reliable *a priori* information using boxplus operation [40] on both *extrinsic* information corresponded to \mathbf{u}_A and \mathbf{u}_R . However, for the scenario $(\Gamma_{AD}, \Gamma_{BD}, p_R) = (-30 \text{ dB}, -30 \text{ dB}, 0)$ with the same assumption, only the soft channel values of \mathbf{y}_{RD} are reliable, and thus the recoveries of \mathbf{u}_A and \mathbf{u}_B require *a priori* information output from boxplus operation. Nevertheless, both *extrinsic* information related to \mathbf{u}_A and \mathbf{u}_B is unreliable with high probabilities, and hence, the output from boxplus operation is still unreliable for the decoding of \mathbf{y}_{iD} for $i \in \{A, B\}$. Therefore, as shown in Fig. 7(c), the value of the probability cannot be reduced to a very small value, although both intra links and RD link are perfect (i.e., $p_A = p_B = p_R = 0$).

VI. CONCLUSION

We have derived the outage probability for the e-MARC system. The theoretical outage probabilities of MARC-NC [8] and MARC-SDF [24] were also included for comparisons in Symmetric/Asymmetric scenarios listed in Table II. It has been observed through simulations that the outage probability of MARC-NC is inferior to that of e-MARC. Moreover, it has been found through simulations that in Asymmetric scenario where one of the source nodes is far away from both the relay and the destination, e-MARC performs better than MARC-SDF. However, in Symmetric scenario, e-MARC is inferior to MARC-SDF.

We have further applied the derivation process of the outage probabilities to two special cases: 1) static intra links and 2) the previously proposed DDEX e-MARC system [16]. According to the numerical results shown in case 1), we verified that erroneous estimates received at the relay is useful for the recovery of the information sequences at the destination. To evaluate the DDEX e-MARC system, comparisons are provided between the e-MARC outage probability and the outage probability of the DDEX e-MARC system in case 2). It was found that a 3.5 to 4.5-dB loss is in exchange for a significant reduction in computational complexity. Finally, we recognized that for the case two source nodes are highly correlated, the source correlation can be exploited to improve the e-MARC outage performance even if only one of the direct links is reliable.

ACKNOWLEDGMENT

The authors would like to thank the reviewers, Mr. Xin He and Dr. Meng Cheng for providing constructive comments to improve this paper significantly.

APPENDIX A DERIVATION OF (7a)-(7d)

We start the derivation of the results of (7a)-(7d) with the general equivalent system model of e-MARC shown in Fig. 8, where source nodes A and B are assumed to be correlated. Variables e_A, e_B, e_R and e_s are i.i.d. binary random variables with $\Pr(e_A = 1) = p_A$, $\Pr(e_B = 1) = p_B$, $\Pr(e_R = 1) = p_R$ and $\Pr(e_s = 1) = p_s$, respectively. According to the definition, the mutual information $I(\mathbf{u}_R; \hat{\mathbf{u}}_R)$ is

$$I(\mathbf{u}_R; \hat{\mathbf{u}}_R) = H(\hat{\mathbf{u}}_R) - H(\hat{\mathbf{u}}_R | \mathbf{u}_R). \quad (28)$$

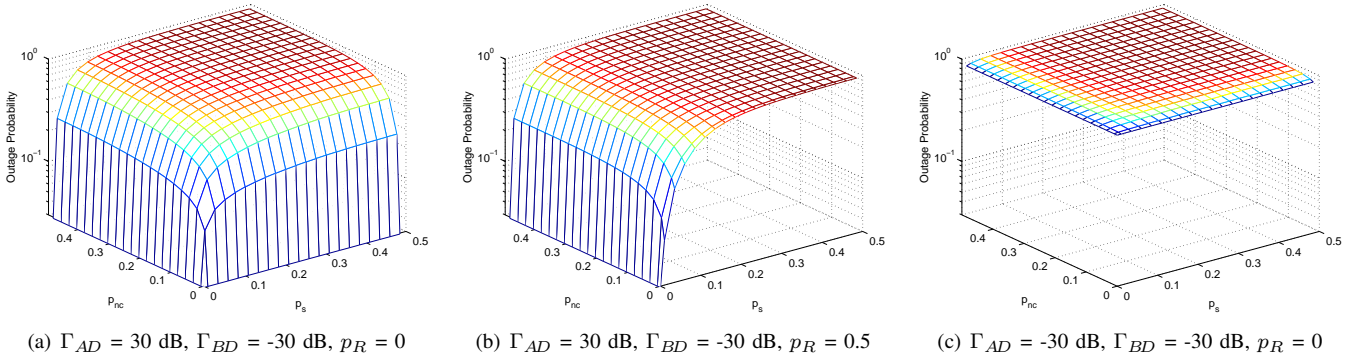


Fig. 7. Impacts of source correlation p_s and network correlation p_{nc} on outage probability of e-MARC. Here $\mathcal{R}_{ci} = \mathcal{R}_{cR} = 1/2$.

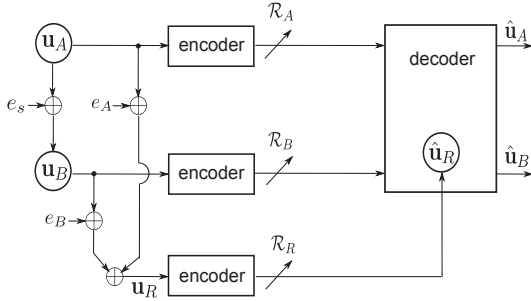


Fig. 8. Generalized equivalent model of e-MARC, where source correlation is taken into account.

Then, given the model shown in Fig. 8, $I(\mathbf{u}_R; \hat{\mathbf{u}}_R)$ can be further expressed as follows

$$\begin{aligned}
 I(\mathbf{u}_R; \hat{\mathbf{u}}_R) &= H(\underbrace{\mathbf{u}_A \oplus e_A}_{\tilde{\mathbf{u}}_A} \oplus \underbrace{\mathbf{u}_B \oplus e_B \oplus e_R}_{\tilde{\mathbf{u}}_B}) - H_b(p_R) \\
 &= H(\underbrace{\mathbf{u}_A \oplus e_A \oplus e_s}_{\tilde{\mathbf{u}}_A} \oplus \underbrace{\mathbf{u}_B \oplus e_B \oplus e_R}_{\tilde{\mathbf{u}}_B}) - H_b(p_R) \\
 &= H(e_A \oplus e_B \oplus e_R \oplus e_s) - H_b(p_R) \\
 &= H_b(p_A * p_B * p_R * p_s) - H_b(p_R), \quad (29)
 \end{aligned}$$

where the operation $\alpha * \beta$ is defined as $\alpha(1 - \beta) + \beta(1 - \alpha)$, and $H_b(\cdot)$ denotes the binary entropy function [29]. For the sake of notation simplicity, we defined μ as $\mu = H(\hat{\mathbf{u}}_R) = H_b(p_A * p_B * p_R * p_s)$. Then, by using the chain rule, the joint entropy $H(\mathbf{u}_A, \mathbf{u}_B, \mathbf{u}_R, \hat{\mathbf{u}}_R)$ can be represented as

$$\begin{aligned}
 H(\mathbf{u}_A, \mathbf{u}_B, \mathbf{u}_R, \hat{\mathbf{u}}_R) &= H(\mathbf{u}_A, \mathbf{u}_B) + H(\mathbf{u}_R | \mathbf{u}_A, \mathbf{u}_B) \\
 &\quad + H(\hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B, \mathbf{u}_R) \quad (30a)
 \end{aligned}$$

$$\begin{aligned}
 &= H(\hat{\mathbf{u}}_R) + H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R) \\
 &\quad + H(\mathbf{u}_R | \mathbf{u}_A, \mathbf{u}_B, \hat{\mathbf{u}}_R). \quad (30b)
 \end{aligned}$$

Hence, the conditional entropy $H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R)$ is expressed with the combination of (30a) and (30b), as

$$\begin{aligned}
 H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R) &= H(\mathbf{u}_A, \mathbf{u}_B) + H(\mathbf{u}_R | \mathbf{u}_A, \mathbf{u}_B) \\
 &\quad + H(\hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B, \mathbf{u}_R) - H(\hat{\mathbf{u}}_R) \\
 &\quad - H(\mathbf{u}_R | \mathbf{u}_A, \mathbf{u}_B, \hat{\mathbf{u}}_R). \quad (31)
 \end{aligned}$$

Also, the conditional mutual information $I(\mathbf{u}_R; \hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B)$ can be expressed by the chain rule, as

$$\begin{aligned}
 I(\mathbf{u}_R; \hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B) &= H(\mathbf{u}_R | \mathbf{u}_A, \mathbf{u}_B) - H(\mathbf{u}_R | \mathbf{u}_A, \mathbf{u}_B, \hat{\mathbf{u}}_R) \\
 &= H(\hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B) - H(\hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B, \mathbf{u}_R). \quad (32)
 \end{aligned}$$

By substituting $H(\mathbf{u}_R | \mathbf{u}_A, \mathbf{u}_B)$ of (32) into (31), $H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R)$ can be rewritten as

$$\begin{aligned}
 H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R) &= H(\mathbf{u}_A, \mathbf{u}_B) + H(\hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B) - H(\hat{\mathbf{u}}_R) \quad (33)
 \end{aligned}$$

where the conditional entropy $H(\hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B)$ is calculated as

$$\begin{aligned}
 H(\hat{\mathbf{u}}_R | \mathbf{u}_A, \mathbf{u}_B) &= H(\mathbf{u}_A \oplus e_A \oplus \mathbf{u}_B \oplus e_B \oplus e_R | \mathbf{u}_A, \mathbf{u}_B) \\
 &= H(e_A \oplus e_B \oplus e_R). \quad (34)
 \end{aligned}$$

Therefore, $H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R)$ shown in the third inequality of (6) can be obtained by substituting (34) into (33), as

$$\begin{aligned}
 H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R) &= H(\mathbf{u}_A, \mathbf{u}_B) + H(e_A \oplus e_B \oplus e_R) - H(\hat{\mathbf{u}}_R) \\
 &= 1 + H_b(p_s) + H_b(p_A * p_B * p_R) - \mu. \quad (35)
 \end{aligned}$$

Again, let $\delta_s = H_b(p_s) + H_b(p_A * p_B * p_R) - \mu$ for simplicity. Since the conditional entropy $H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R)$ in (35) can also alternatively be expressed by the chain rule, as

$$H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R) = H(\mathbf{u}_B | \mathbf{u}_A, \hat{\mathbf{u}}_R) + H(\mathbf{u}_A | \hat{\mathbf{u}}_R) \quad (36a)$$

$$= H(\mathbf{u}_A | \mathbf{u}_B, \hat{\mathbf{u}}_R) + H(\mathbf{u}_B | \hat{\mathbf{u}}_R). \quad (36b)$$

Since the sequence $\hat{\mathbf{u}}_R$ is composed of \mathbf{u}_A , the conditional entropy $H(\mathbf{u}_A | \hat{\mathbf{u}}_R)$ in (36a) can be calculated as

$$\begin{aligned}
 H(\mathbf{u}_A | \hat{\mathbf{u}}_R) &= H(\mathbf{u}_A | \mathbf{u}_A \oplus e_A \oplus \mathbf{u}_B \oplus e_B \oplus e_R) \\
 &= H(e_A \oplus \mathbf{u}_A \oplus e_s \oplus e_B \oplus e_R) \\
 &= 1. \quad (37)
 \end{aligned}$$

Therefore, we can obtain $H(\mathbf{u}_B | \mathbf{u}_A, \hat{\mathbf{u}}_R)$, shown in the second inequality of (6), by substituting (35) and (37) into (36a), as

$$\begin{aligned}
 H(\mathbf{u}_B | \mathbf{u}_A, \hat{\mathbf{u}}_R) &= \underbrace{H(\mathbf{u}_A, \mathbf{u}_B | \hat{\mathbf{u}}_R)}_{1+\delta_s} - \underbrace{H(\mathbf{u}_A | \hat{\mathbf{u}}_R)}_1 \\
 &= \delta_s. \quad (38)
 \end{aligned}$$

The conditional entropy $H(\mathbf{u}_A|\mathbf{u}_B, \hat{\mathbf{u}}_R)$ shown in the first inequality of (6) is also calculated in the same way, and the result is

$$H(\mathbf{u}_A|\mathbf{u}_B, \hat{\mathbf{u}}_R) = \delta_s. \quad (39)$$

A set of inequalities is obtained by combining (29), (35), (38) and (39), as

$$\begin{aligned} \mathcal{R}_A &\geq \delta_s, \\ \mathcal{R}_B &\geq \delta_s, \\ \mathcal{R}_A + \mathcal{R}_B &\geq 1 + \delta_s, \\ \mathcal{R}_R &\geq \mu - H_b(p_R). \end{aligned} \quad (40)$$

For the case that the two source nodes are uncorrelated (i.e., $p_s = 0.5$), $\mu = 1$ and $\delta_s = \delta = H_b(p_A * p_B * p_R)$, which leads (40) into (7a)-(7d).

APPENDIX B

DIVISIONS OF INTEGRALS (21) AND (22)

As shown in (16) and (18), the values of $\hat{p}_i(\gamma_{iR})$ and $\hat{p}_R(\gamma_{RD})$ become zero as the SNR γ_{iR} and γ_{RD} excess over the threshold value, and thus, the domain of the integral shown in (23) can be divided into eight intervals according to the threshold value γ_A^* , γ_B^* and γ_R^* , as

$$\begin{aligned} \Pr(\varepsilon_1) &= \mathfrak{J}[g_1; V] = \sum_{j=1}^8 \mathfrak{J}[g_1; V_j] \\ \Pr(\varepsilon_2) &= \mathfrak{J}[g_2; V] = \sum_{j=1}^8 \mathfrak{J}[g_2; V_j], \end{aligned} \quad (41)$$

where

$$\begin{aligned} V_1 &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{A}^+, \gamma_{BR} \in \mathbb{B}^+, \gamma_{RD} \in \mathbb{C}^+\} \\ V_2 &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{A}^+, \gamma_{BR} \in \mathbb{B}^-, \gamma_{RD} \in \mathbb{C}^+\} \\ V_3 &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{A}^-, \gamma_{BR} \in \mathbb{B}^+, \gamma_{RD} \in \mathbb{C}^+\} \\ V_4 &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{A}^-, \gamma_{BR} \in \mathbb{B}^-, \gamma_{RD} \in \mathbb{C}^+\} \\ V_5 &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{A}^+, \gamma_{BR} \in \mathbb{B}^+, \gamma_{RD} \in \mathbb{C}^-\} \\ V_6 &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{A}^+, \gamma_{BR} \in \mathbb{B}^-, \gamma_{RD} \in \mathbb{C}^-\} \\ V_7 &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{A}^-, \gamma_{BR} \in \mathbb{B}^+, \gamma_{RD} \in \mathbb{C}^-\} \\ V_8 &= \{(\gamma_{AR}, \gamma_{BR}, \gamma_{RD}) : \gamma_{AR} \in \mathbb{A}^-, \gamma_{BR} \in \mathbb{B}^-, \gamma_{RD} \in \mathbb{C}^-\}. \end{aligned} \quad (42)$$

Here, the sets \mathbb{A}^- , \mathbb{A}^+ , \mathbb{B}^- , \mathbb{B}^+ , \mathbb{C}^- and \mathbb{C}^+ are defined as

$$\begin{aligned} \mathbb{A}^- &= [0, \gamma_A^*), & \mathbb{A}^+ &= [\gamma_A^*, \infty) \\ \mathbb{B}^- &= [0, \gamma_B^*), & \mathbb{B}^+ &= [\gamma_B^*, \infty) \\ \mathbb{C}^- &= [0, \gamma_R^*), & \mathbb{C}^+ &= [\gamma_R^*, \infty). \end{aligned} \quad (43)$$

The integrals of (21) and (22) with respect to each interval listed in (42) are expressed in (44) and (45), respectively. The arguments $\hat{\delta}_j$ and $\hat{\omega}_j$ corresponding to each interval in (44) and (45) are summarized in Table IV.

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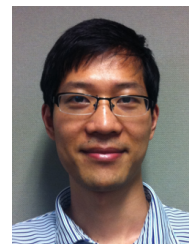
TABLE IV
ARGUMENTS OF FUNCTIONS $f_A^{-1}(\cdot)$ AND $f_B^{-1}(\cdot)$ IN (44) AND (45).

$\hat{\delta}_1$	$H_b(0 * 0 * 0)$	$\hat{\omega}_1$	$\hat{\delta}_1 + 1 - f_A(\gamma_{AD})$
$\hat{\delta}_2$	$H_b(0 * \tilde{p}_B(\gamma_{BR}) * 0)$	$\hat{\omega}_2$	$\hat{\delta}_2 + 1 - f_A(\gamma_{AD})$
$\hat{\delta}_3$	$H_b(\tilde{p}_A(\gamma_{AR}) * 0 * 0)$	$\hat{\omega}_3$	$\hat{\delta}_3 + 1 - f_A(\gamma_{AD})$
$\hat{\delta}_4$	$H_b(\tilde{p}_A(\gamma_{AR}) * \tilde{p}_B(\gamma_{BR}) * 0)$	$\hat{\omega}_4$	$\hat{\delta}_4 + 1 - f_A(\gamma_{AD})$
$\hat{\delta}_5$	$H_b(0 * 0 * \tilde{p}_R(\gamma_{RD}))$	$\hat{\omega}_5$	$\hat{\delta}_5 + 1 - f_A(\gamma_{AD})$
$\hat{\delta}_6$	$H_b(0 * \tilde{p}_B(\gamma_{BR}) * \tilde{p}_R(\gamma_{RD}))$	$\hat{\omega}_6$	$\hat{\delta}_6 + 1 - f_A(\gamma_{AD})$
$\hat{\delta}_7$	$H_b(\tilde{p}_A(\gamma_{AR}) * 0 * \tilde{p}_R(\gamma_{RD}))$	$\hat{\omega}_7$	$\hat{\delta}_7 + 1 - f_A(\gamma_{AD})$
$\hat{\delta}_8$	$H_b(\tilde{p}_A(\gamma_{AR}) * \tilde{p}_B(\gamma_{BR}) * \tilde{p}_R(\gamma_{RD}))$	$\hat{\omega}_8$	$\hat{\delta}_8 + 1 - f_A(\gamma_{AD})$

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$$\begin{aligned}
\mathcal{J}[g_1; V_1] &= \frac{1}{\Gamma_{AD}} \exp\left(-\frac{\gamma_R^*}{\Gamma_{RD}} - \frac{\gamma_A^*}{\Gamma_{AR}} - \frac{\gamma_B^*}{\Gamma_{BR}}\right) \int_{f_A^{-1}(\hat{\delta}_1)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega}_1)}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}}\right) d\gamma_{AD} \\
\mathcal{J}[g_1; V_2] &= \frac{1}{\Gamma_{AD}\Gamma_{BR}} \exp\left(-\frac{\gamma_R^*}{\Gamma_{RD}} - \frac{\gamma_A^*}{\Gamma_{AR}}\right) \int_0^{\gamma_B^*} \int_{f_A^{-1}(\hat{\delta}_2)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega}_2)}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}} - \frac{\gamma_{BR}}{\Gamma_{BR}}\right) d\gamma_{AD} d\gamma_{BR} \\
\mathcal{J}[g_1; V_3] &= \frac{1}{\Gamma_{AD}\Gamma_{AR}} \exp\left(-\frac{\gamma_R^*}{\Gamma_{RD}} - \frac{\gamma_B^*}{\Gamma_{BR}}\right) \int_0^{\gamma_A^*} \int_{f_A^{-1}(\hat{\delta}_3)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega}_3)}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}} - \frac{\gamma_{AR}}{\Gamma_{AR}}\right) d\gamma_{AD} d\gamma_{AR} \\
\mathcal{J}[g_1; V_4] &= \frac{1}{\Gamma_{AD}\Gamma_{AR}\Gamma_{BR}} \exp\left(-\frac{\gamma_R^*}{\Gamma_{RD}}\right) \int_0^{\gamma_B^*} \int_0^{\gamma_A^*} \int_{f_A^{-1}(\hat{\delta}_4)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega}_4)}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}} - \frac{\gamma_{AR}}{\Gamma_{AR}} - \frac{\gamma_{BR}}{\Gamma_{BR}}\right) d\gamma_{AD} d\gamma_{AR} d\gamma_{BR} \\
\mathcal{J}[g_1; V_5] &= \frac{1}{\Gamma_{AD}\Gamma_{RD}} \exp\left(-\frac{\gamma_A^*}{\Gamma_{AR}} - \frac{\gamma_B^*}{\Gamma_{BR}}\right) \int_0^{\gamma_R^*} \int_{f_A^{-1}(\hat{\delta}_5)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega}_5)}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}} - \frac{\gamma_{RD}}{\Gamma_{RD}}\right) d\gamma_{AD} d\gamma_{RD} \\
\mathcal{J}[g_1; V_6] &= \frac{1}{\Gamma_{AD}\Gamma_{RD}\Gamma_{BR}} \exp\left(-\frac{\gamma_A^*}{\Gamma_{AR}}\right) \int_0^{\gamma_R^*} \int_0^{\gamma_B^*} \int_{f_A^{-1}(\hat{\delta}_6)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega}_6)}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}} - \frac{\gamma_{BR}}{\Gamma_{BR}} - \frac{\gamma_{RD}}{\Gamma_{RD}}\right) d\gamma_{AD} d\gamma_{BR} d\gamma_{RD} \\
\mathcal{J}[g_1; V_7] &= \frac{1}{\Gamma_{AD}\Gamma_{RD}\Gamma_{AR}} \exp\left(-\frac{\gamma_B^*}{\Gamma_{BR}}\right) \int_0^{\gamma_R^*} \int_0^{\gamma_A^*} \int_{f_A^{-1}(\hat{\delta}_7)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega}_7)}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}} - \frac{\gamma_{AR}}{\Gamma_{AR}} - \frac{\gamma_{RD}}{\Gamma_{RD}}\right) d\gamma_{AD} d\gamma_{AR} d\gamma_{RD} \\
\mathcal{J}[g_1; V_8] &= \frac{1}{\Gamma_{AD}\Gamma_{RD}\Gamma_{AR}\Gamma_{BR}} \int_0^{\gamma_R^*} \int_0^{\gamma_B^*} \int_0^{\gamma_A^*} \int_{f_A^{-1}(\hat{\delta}_8)}^{f_A^{-1}(1)} \exp\left(\frac{-f_B^{-1}(\hat{\omega}_8)}{\Gamma_{BD}} - \frac{\gamma_{AD}}{\Gamma_{AD}} - \frac{\gamma_{AR}}{\Gamma_{AR}} - \frac{\gamma_{BR}}{\Gamma_{BR}} - \frac{\gamma_{RD}}{\Gamma_{RD}}\right) d\gamma_{AD} d\gamma_{AR} d\gamma_{BR} d\gamma_{RD}
\end{aligned} \tag{44}$$

$$\begin{aligned}
\mathcal{J}[g_2; V_1] &= \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{\gamma_R^*}{\Gamma_{RD}} - \frac{\gamma_A^*}{\Gamma_{AR}} - \frac{\gamma_B^*}{\Gamma_{BR}}\right) \\
\mathcal{J}[g_2; V_2] &= \frac{1}{\Gamma_{BR}} \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{\gamma_R^*}{\Gamma_{RD}} - \frac{\gamma_A^*}{\Gamma_{AR}}\right) \int_0^{\gamma_B^*} \exp\left(\frac{-f_B^{-1}(\hat{\delta}_2)}{\Gamma_{BD}} - \frac{\gamma_{BR}}{\Gamma_{BR}}\right) d\gamma_{BR} \\
\mathcal{J}[g_2; V_3] &= \frac{1}{\Gamma_{AR}} \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{\gamma_R^*}{\Gamma_{RD}} - \frac{\gamma_B^*}{\Gamma_{BR}}\right) \int_0^{\gamma_A^*} \exp\left(\frac{-f_B^{-1}(\hat{\delta}_3)}{\Gamma_{BD}} - \frac{\gamma_{AR}}{\Gamma_{AR}}\right) d\gamma_{AR} \\
\mathcal{J}[g_2; V_4] &= \frac{1}{\Gamma_{AR}\Gamma_{BR}} \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{\gamma_R^*}{\Gamma_{RD}}\right) \int_0^{\gamma_B^*} \int_0^{\gamma_A^*} \exp\left(\frac{-f_B^{-1}(\hat{\delta}_4)}{\Gamma_{BD}} - \frac{\gamma_{AR}}{\Gamma_{AR}} - \frac{\gamma_{BR}}{\Gamma_{BR}}\right) d\gamma_{AR} d\gamma_{BR} \\
\mathcal{J}[g_2; V_5] &= \frac{1}{\Gamma_{RD}} \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{\gamma_A^*}{\Gamma_{AR}} - \frac{\gamma_B^*}{\Gamma_{BR}}\right) \int_0^{\gamma_R^*} \exp\left(\frac{-f_B^{-1}(\hat{\delta}_5)}{\Gamma_{BD}} - \frac{\gamma_{RD}}{\Gamma_{RD}}\right) d\gamma_{RD} \\
\mathcal{J}[g_2; V_6] &= \frac{1}{\Gamma_{RD}\Gamma_{BR}} \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{\gamma_A^*}{\Gamma_{AR}}\right) \int_0^{\gamma_R^*} \int_0^{\gamma_B^*} \exp\left(\frac{-f_B^{-1}(\hat{\delta}_6)}{\Gamma_{BD}} - \frac{\gamma_{BR}}{\Gamma_{BR}} - \frac{\gamma_{RD}}{\Gamma_{RD}}\right) d\gamma_{BR} d\gamma_{RD} \\
\mathcal{J}[g_2; V_7] &= \frac{1}{\Gamma_{RD}\Gamma_{AR}} \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}} - \frac{\gamma_B^*}{\Gamma_{BR}}\right) \int_0^{\gamma_R^*} \int_0^{\gamma_A^*} \exp\left(\frac{-f_B^{-1}(\hat{\delta}_7)}{\Gamma_{BD}} - \frac{\gamma_{AR}}{\Gamma_{AR}} - \frac{\gamma_{RD}}{\Gamma_{RD}}\right) d\gamma_{AR} d\gamma_{RD} \\
\mathcal{J}[g_2; V_8] &= \frac{1}{\Gamma_{RD}\Gamma_{AR}\Gamma_{BR}} \exp\left(\frac{-f_A^{-1}(1)}{\Gamma_{AD}}\right) \int_0^{\gamma_R^*} \int_0^{\gamma_B^*} \int_0^{\gamma_A^*} \exp\left(\frac{-f_B^{-1}(\hat{\delta}_8)}{\Gamma_{BD}} - \frac{\gamma_{AR}}{\Gamma_{AR}} - \frac{\gamma_{BR}}{\Gamma_{BR}} - \frac{\gamma_{RD}}{\Gamma_{RD}}\right) d\gamma_{AR} d\gamma_{BR} d\gamma_{RD}
\end{aligned} \tag{45}$$



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